

Appendix 3

Interrelation between ΔS^{shuf} and σ/μ in the case of IID

If we consider Q_k , where $Q_k \geq 0, k = 1, 2, \dots, N$, we obtain the quantities

$$p_k = Q_k / \sum_{l=1}^N Q_l,$$

which satisfy the necessary conditions[1]: $p_k \geq 0, \sum_{k=1}^N p_k = 1$ to be considered as point probabilities. We then define[2–4] the moments of the natural time $\chi_k = k/N$ as

$$\langle \chi^q \rangle = \sum_{k=1}^N (k/N)^q p_k$$

and the entropy

$$S \equiv \langle \chi \ln \chi \rangle - \langle \chi \rangle \ln \langle \chi \rangle,$$

where

$$\langle \chi \ln \chi \rangle = \sum_{k=1}^N (k/N) \ln(k/N) p_k.$$

The effect of the time reversal operator[5] on Q_k is obtained by $\mathcal{T}p_k = p_{N-k+1}$, and, similarly, the time-reversed entropy $\mathcal{T}S (\equiv S_-)$ is obtained by the same formula as S but by using p_{N-k+1} instead of p_k .

The relevant expressions for evaluating S and S_- , when a window of length $i (= N)$ is sliding pulse by pulse through a time series, are given in Appendix 2 deposited in the same directory as this file.

Here, we consider the case when Q_k are independent and identically distributed (IID) positive random variables. It then follows that the expectation value $E(p_k) = E[Q_k / \sum_{l=1}^N Q_l]$ of p_k equals $1/N$:

$$E(p_k) = \frac{1}{N}. \tag{1}$$

Equation (1) results from the fact that, since Q_k are IID, we have:

$$E\left[\sum_{k=1}^N Q_k / \sum_{l=1}^N Q_l\right] = 1 = NE(p_k).$$

For the purpose of our calculations, the relation between the variance of p_k ,

$$\text{Var}(p_k) = E[(p_k - 1/N)^2],$$

and the covariance of p_k and p_l ,

$$\text{Cov}(p_k, p_l) = \text{E}[(p_k - 1/N)(p_l - 1/N)],$$

is of central importance. Using the constraint $\sum_{k=1}^N p_k = 1$, leading to $p_k - 1/N = \sum_{l \neq k} (1/N - p_l)$, and the fact that Q_k are IID, we obtain

$$\text{E}[(p_k - 1/N)^2] = \text{E} \left[(p_k - 1/N) \sum_{l \neq k} (1/N - p_l) \right] = -(N-1) \text{E}[(p_k - 1/N)(p_l - 1/N)].$$

Thus, we get

$$\text{Cov}(p_k, p_l) = -\frac{\text{Var}(p_k)}{N-1}. \quad (2)$$

The N -dependence of $\text{Var}(p_k)$ is obtained from

$$\text{Var}(p_k) = \frac{1}{N^2} \text{E} \left[\left(\frac{NQ_k}{\sum_{l=1}^N Q_l} - 1 \right)^2 \right], \quad (3)$$

where the quantity $\text{E}[(NQ_k/\sum_{l=1}^N Q_l - 1)^2]$ is *asymptotically* N -independent. The latter arises as follows: If $\text{E}(Q_k) = \mu$ and $\text{Var}(Q_k) = \sigma^2 (< \infty)$, as a result of the central limit theorem[6], we have $\text{E}(\sum_{k=1}^N Q_k/N) = \mu$ and $\text{Var}(\sum_{k=1}^N Q_k/N) = \sigma^2/N$. The latter two equations, for large enough N imply that

$$\text{E}[(NQ_k/\sum_{l=1}^N Q_l - 1)^2] \approx \text{E}[(Q_k/\mu - 1)^2] = \sigma^2/\mu^2.$$

Thus, Eq.(3) becomes

$$\text{Var}(p_k) = \frac{\sigma^2}{N^2\mu^2}. \quad (4)$$

Let us first study $\text{E}[S - \mathcal{T}S] = \text{E}(S) - \text{E}(\mathcal{T}S)$ for which we intuitively expect that it equals zero for positive IID Q_k . Indeed, we have that

$$\text{E}(S) = \text{E} \left[\sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{N} \right) p_k - \sum_{k=1}^N \frac{k}{N} p_k \ln \left(\sum_{l=1}^N \frac{l}{N} p_l \right) \right] \quad (5)$$

and

$$\text{E}(\mathcal{T}S) = \text{E} \left[\sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{N} \right) p_{N-k+1} - \sum_{k=1}^N \frac{k}{N} p_{N-k+1} \ln \left(\sum_{l=1}^N \frac{l}{N} p_{N-l+1} \right) \right]. \quad (6)$$

The result of Eq.(5) depends on $\text{Var}(p_k)$ and $\text{Cov}(p_k, p_l)$ (see Ref.[7]), whereas that of Eq.(6) on $\text{Var}(p_{N-k+1})$ and $\text{Cov}(p_{N-k+1}, p_{N-l+1})$. In view of the fact that both $\text{Var}(p_k)$ and $\text{Cov}(p_k, p_l)$ are independent of k and l (see Eqs.(2) and (4) above), we have

$$\text{E}(\mathcal{T}S) = \text{E}(S)$$

and moreover[7]:

$$E(S) = E(TS) = \sum_{k=1}^N \frac{k}{N^2} \ln \left(\frac{k}{N\bar{\chi}} \right) - \frac{\sigma^2}{(N-1)\mu^2} \left(\frac{\bar{\chi}^2}{\bar{\chi}} - \bar{\chi} \right), \quad (7)$$

where $\bar{\chi} = \sum k/N^2 = (1 + 1/N)/2$ and $\bar{\chi}^2 = \sum k^2/N^3 = (1 + 1/N)[1 + 1/(2N)]/3$.

We now turn to the variance of $\Delta S \equiv S - TS$ defined by

$$\sigma^2[\Delta S] \equiv E \{ [S - TS - E(S - TS)]^2 \} = E [(S - TS)^2],$$

in view of Eq.(7), which is of primary importance in ECG. We have that

$$\begin{aligned} \sigma^2[\Delta S] &= E [(S - TS)^2] = E \{ [S - E(S)] - [TS - E(TS)] \}^2 \\ &= E \{ [S - E(S)]^2 \} + E \{ [TS - E(TS)]^2 \} - 2E \{ [S - E(S)] [TS - E(TS)] \} \\ &= 2 [\delta S^2 - E \{ [S - E(S)] [TS - E(TS)] \}], \end{aligned} \quad (8)$$

where we used the fact that $\delta S^2 \equiv E \{ [S - E(S)]^2 \}$ (originally defined in Ref.[8], see also Ref.[7]) remains unchanged under time reversal for the same reasons as $E(S) = E(TS)$. The term $E \{ [S - E(S)] [TS - E(TS)] \}$ can be evaluated in a way similar to the one used in Ref.[7]. Namely, we add and subtract the term $\sum_{k=1}^N \frac{k}{N} p_k \ln \bar{\chi}$ from S and the term $\sum_{k=1}^N \frac{k}{N} p_{N-k+1} \ln \bar{\chi}$ from TS . We then expand the resulting logarithmic terms $\ln[1 + \sum_{l=1}^N \frac{l}{N} (p_l - \frac{1}{N})/\bar{\chi}]$ and $\ln[1 + \sum_{l=1}^N \frac{l}{N} (p_{N-l+1} - \frac{1}{N})/\bar{\chi}]$ to first order in $(p_l - \frac{1}{N})$ and $(p_{N-l+1} - \frac{1}{N})$, respectively. This leads to

$$\begin{aligned} E \{ [S - E(S)] [TS - E(TS)] \} &= E \left\{ \left[\sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{eN\bar{\chi}} \right) \left(p_k - \frac{1}{N} \right) - \right. \right. \\ &\quad \left. \left. - \frac{1}{\bar{\chi}} \sum_{k=1}^N \frac{k}{N} \left(p_k - \frac{1}{N} \right) \sum_{l=1}^N \frac{l}{N} \left(p_l - \frac{1}{N} \right) + \frac{\sigma^2}{(N-1)\mu^2} \left(\frac{\bar{\chi}^2}{\bar{\chi}} - \bar{\chi} \right) \right] \times \right. \\ &\quad \times \left[\sum_{k'=1}^N \frac{k'}{N} \ln \left(\frac{k'}{eN\bar{\chi}} \right) \left(p_{N-k'+1} - \frac{1}{N} \right) - \right. \\ &\quad \left. \left. - \frac{1}{\bar{\chi}} \sum_{k'=1}^N \frac{k'}{N} \left(p_{N-k'+1} - \frac{1}{N} \right) \sum_{l'=1}^N \frac{l'}{N} \left(p_{N-l'+1} - \frac{1}{N} \right) + \right. \\ &\quad \left. \left. + \frac{\sigma^2}{(N-1)\mu^2} \left(\frac{\bar{\chi}^2}{\bar{\chi}} - \bar{\chi} \right) \right] \right\}, \end{aligned} \quad (9)$$

where $e = 2.7182\dots$ is the base of natural logarithms. *If we assume that the distribution of*

Q_k is symmetric around μ and keeping up to the first order in σ^2/μ^2 , Eq.(9) simplifies to

$$\begin{aligned} \mathbb{E} \{[S - \mathbb{E}(S)] [\mathcal{T}S - \mathbb{E}(\mathcal{T}S)]\} &= \mathbb{E} \left\{ \left[\sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{eN\bar{\chi}} \right) \left(p_k - \frac{1}{N} \right) \right] \times \right. \\ &\quad \left. \times \left[\sum_{k'=1}^N \frac{k'}{N} \ln \left(\frac{k'}{eN\bar{\chi}} \right) \left(p_{N-k'+1} - \frac{1}{N} \right) \right] \right\}. \end{aligned} \quad (10)$$

Now, using Eqs.(2) and (4), we obtain

$$\mathbb{E} \left\{ \left(p_k - \frac{1}{N} \right) \left(p_{N-k'+1} - \frac{1}{N} \right) \right\} = \frac{\sigma^2}{(N-1)N\mu^2} \delta_{k,N-k'+1} - \frac{\sigma^2}{(N-1)N^2\mu^2}, \quad (11)$$

where $\delta_{l,m}$ is Kronecker’s delta (equal to 1 if $l = m$, and 0 otherwise). Substituting Eq.(11) into Eq.(10), we finally find that

$$\begin{aligned} \mathbb{E} \{[S - \mathbb{E}(S)] [\mathcal{T}S - \mathbb{E}(\mathcal{T}S)]\} &= \frac{\sigma^2}{(N-1)\mu^2} \left\{ \sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{eN\bar{\chi}} \right) \times \right. \\ &\quad \times \frac{N-k+1}{N} \ln \left(\frac{N-k+1}{eN\bar{\chi}} \right) \frac{1}{N} - \\ &\quad \left. - \left[\sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{eN\bar{\chi}} \right) \frac{1}{N} \right]^2 \right\} \end{aligned} \quad (12)$$

Using now Eq.(A21) of Ref.[7] for δS^2 , i.e.,

$$\delta S^2 = \frac{\sigma^2}{(N-1)\mu^2} \left\{ \sum_{k=1}^N \left(\frac{k}{N} \ln \frac{k}{eN\bar{\chi}} \right)^2 \frac{1}{N} - \left[\sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{eN\bar{\chi}} \right) \frac{1}{N} \right]^2 \right\}, \quad (13)$$

we obtain

$$\begin{aligned} \sigma^2[\Delta S] &= \frac{2\sigma^2}{(N-1)\mu^2} \left[\sum_{k=1}^N \left(\frac{k}{N} \ln \frac{k}{eN\bar{\chi}} \right)^2 \frac{1}{N} - \right. \\ &\quad \left. - \sum_{k=1}^N \frac{k}{N} \ln \left(\frac{k}{eN\bar{\chi}} \right) \frac{N-k+1}{N} \ln \left(\frac{N-k+1}{eN\bar{\chi}} \right) \frac{1}{N} \right] \end{aligned} \quad (14)$$

Equation (14) reflects that, when a window of length $i(= N)$ is sliding through the randomly shuffled Q_k of an ECG, the following relation holds

$$\sigma[\Delta S_i^{shuf}] = \frac{\sigma}{\mu} \sqrt{f(i)}, \quad (15)$$

where

$$f(i) = \frac{2}{i-1} \left[\sum_{k=1}^i \left(\frac{k}{i} \ln \frac{k}{ie\bar{\chi}} \right)^2 \frac{1}{i} - \sum_{k=1}^i \frac{k}{i} \ln \left(\frac{k}{ie\bar{\chi}} \right) \frac{i-k+1}{i} \ln \left(\frac{i-k+1}{ie\bar{\chi}} \right) \frac{1}{i} \right].$$

Additional information for the paper ‘*Identifying sudden cardiac death risk and specifying its occurrence time by analyzing electrocardiograms in natural time*’

In the main text, the numerator in the measure $N_i \equiv \sigma[\Delta S_i^{shuf}]/\sigma[\Delta S_i]$ was calculated on the basis of Eq.(15).

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